# 1 Russian winter and spring wheat productivity, heat stress and

# drought conditions at flowering, and the role of atmospheric

# 3 blocking

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Abstract. Russia has become the foremost wheat exporting country worldwide. Episodic production breakdowns caused by extreme weather during sensitive stages of crop development are hence of concern not only for the domestic but also for the global wheat market. In this study, we examine heat stress occurrence and availability of rainwater during the flowering period, and investigate their impacts on the winter and spring wheat yields of three major production areas of Russia. We also consider the role of atmospheric blocking as a precursor of extreme weather and assess the correlation between blocking duration and yield. Owing to the later occurrence of flowering in spring wheat and the warmer climate of Southern European Russia, we find the probability of heat stress to be higher in spring than in winter wheat, and higher in the south than in the north of the study area. For spring wheat, the negative association between yield and heat stress is stronger than the positive association between yield and total precipitation. The reverse is true for winter wheat. In all regions and for both wheat types, heat stress occurrence and total precipitation amounts correlate significantly with the area-weighted average of blocked days. We also find correlation between blocking duration and yield, but results are significant only for spring wheat.

25 KEY WORDS: wheat, atmospheric blocking, heat stress, drought, Russia

#### 1 Introduction

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The Russian Federation considerably increased its wheat exports during the recent past, 27 becoming in 2016 the leading wheat exporter worldwide (United States Department of 28 Agriculture 2016). Episodic production collapses in Russia can therefore have considerable 29 repercussions on the global wheat market (Zampieri et al. 2017). This occurred in 2010, when 30 a large fraction of the wheat cultivation area of Russia, experienced extraordinary high 31 temperatures throughout the summer (Grumm 2011; Barriopedro et al. 2011; Dole et al. 2011; 32 Wright et al. 2014; Katsafados et al. 2014; Russo et al. 2015). Consequently, harvested grains 33 dropped off by about 40% from previous years' levels, whereupon the Russian government 34 issued a wheat export ban that lasted until mid of 2011 and led to an upsurge in global prices 35 of up to 50 % (Food and Agriculture Organization of the United Nations 2010; Welton 2011). 36 37 Exposure to high temperatures is recurrent in the main wheat production area in the south-west 38 of Russia. Since most of the cultivation is rainfed, lack of precipitation can also limit production 39 (Alcamo et al. 2007; Daryanto et al. 2016). Not all development stages of wheat are very 40 sensitive to water shortage (Brouwer et al. 1989), but water shortage occurring in the time from 41 booting to early grain formation can depress yields (Schneider et al. 1969; Mogensen et al. 42 1985). Even more harmful than drought alone is the co-occurrence of heat stress and drought 43 (Nicolas et al. 1984), a situation that was responsible for the severe production losses suffered 44 in 2010 (Lupo et al. 2014). 45 46 Heat stress, often in association with water shortage, is particularly harmful to wheat if it occurs 47 at flowering (Brouwer et al. 1989; Wheeler et al. 2000; Wollenweber et al. 2003). Production 48 losses caused by temperatures in excess of critical thresholds during this phase of development 49 can indeed be considerable (Fontana et al. 2015). In global investigations of the exposure of 50 crops to damaging temperatures during the reproductive period (Gourdji et al. 2013; Teixeira 51

et al. 2013) some areas within the Russian Federation appear as hot spots. Yet, without addressing the problem at the regional scale, it is not possible to provide detailed information on the recurrence of heat stress and drought conditions. Another limitation of global studies is that they consider wheat as a crop in a generic way without making a distinction between winter and spring varieties, though this is vital in this context, as the two varieties flower asynchronously.

In this paper, we assess the occurrence of high temperature and low precipitation amounts in the reproductive period based on local weather data, and examine their association to winter and spring wheat yield variations at the regional scale based on official yield statistics.

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Atmospheric blocking events are among the large-scale flow patterns that favour the occurrence 62 of extreme heat and drought (Tyrlis & Hoskins 2008; Petoukhov et al. 2013). The presence of 63 long-lasting anticyclones during the growing season of wheat is not unusual in Russia (Park & 64 Ahn 2014; Antokhina et al. 2016). A second objective is therefore to provide a statistical 65 analysis of the influence of atmospheric blocking on the occurrence of heat stress and drought 66 conditions in the reproductive period of wheat and to evaluate in a direct way the correlation 67 between the duration of atmospheric blocking and wheat yield, again considering winter and 68 spring varieties separately. 69

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#### 2 Data and data processing

#### 72 2.1 Daily weather data

The analysis relies on 118 land-based weather stations distributed across the wheat production area of Russia (Figure 1). The database includes daily minimum ( $T_{min}$ ), and maximum temperature ( $T_{max}$ ), and daily precipitation amounts (Pr) in liquid water equivalent. The data were obtained from the Carbon Dioxide Information Analysis Centre (CDIAC) for the period

- 77 1980-2009 (Bulygina & Razuvaev 2012) and the Federal Service of Hydrometeorology and
- 78 Environmental Monitoring (2008) for the period 2010-2014.

- A data quality control was performed using the RClimDex tool (Zhang et al. 2004) to eliminate
- unreasonable values (i.e. Pr < 0 and  $T_{max} < T_{min}$ ). The test also checked for temperature
- values outside a range given by the climatological daily mean  $\pm$  3 times the corresponding
- climatological daily standard deviation, individually for each of  $T_{min}$  and  $T_{max}$ .

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#### 2.2 Blocking Index

- 86 To quantify the occurrence of blocking situations, we adopted the two-dimensional Blocking
- Index (BI) of Rohrer et al. (2018), which is available at  $2^{\circ} \times 2^{\circ}$  spatial resolution and 6-hourly
- 88 temporal resolution. The BI was computed from 6-hourly ERA-interim re-analyses of the 500
- hPa geopotential height (z500) for 1979-2015 (Dee et al. 2011), based on the blocking detection
- 90 algorithm of Lejenäs and Økland (1983), Tibaldi and Molteni (1990) and Tibaldi et al. (1994).
- For a given latitude  $\varphi$  between  $36^{\circ} 76^{\circ}$  North or South, the algorithm checks whether:
- 1) The z500 gradient towards the pole =  $\frac{z500_{\phi+14}-z500_{\phi}}{14} < -10\frac{gpm}{lat}$  and
- 93 2) The z500 gradient towards the equator =  $\frac{z500_{\varphi}-z500_{\varphi-14}}{14} > 0 \frac{gpm}{lat}$
- A blocking condition is detected when areas characterized by a reversal of the meridional z500
- 95 gradient have a spatial overlap of more than 70% during at least five consecutive days
- 96 (Schwierz et al. 2004; Scherrer et al. 2006). The binary BI is the result of assigning a value of
- one when all conditions are satisfied, and of zero otherwise.

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#### 99 2.3 Yield

- 100 Yield statistics for winter (WW) and spring wheat (SW) for the period 1995-2014, as compiled
- at the provincial level by the Federal State Statistics Service of Russia, were obtained from

Schierhorn et al. (2014) along with statistics of the corresponding sowing area for the period 1995-2009. The data were complemented with information available from the Statistical Handbooks of the Russian Statistics Service (Russian Federal State Statistics 2014; Russian Federal State Statistics 2016).

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## 107 **2.4 Wheat phenology**

As in Trnka et al. (2014), wheat phenological development was estimated with the help of the empirical model of Olesen et al. (2012). The model calculates the duration of phenological phases as:

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$$S = \sum \max(T_i - T_b, 0) \cdot \alpha_i$$
 (1)

where S is the temperature sum required to complete the given phase,  $T_i$  is the daily mean temperature,  $T_b$  is the base temperature and  $\alpha_i$  is a daily photoperiodic response. The model assumes that the latter is relevant only in the period from emergence until flowering and gives  $\alpha_i$  as:

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$$\alpha_i = \min\left\{\max\left[\frac{(\lambda_i - 7)}{13}, 0\right], 1\right\} \tag{2}$$

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- where  $\lambda_i$  is the daylength on day of the year i, obtained from latitude using standard astronomical formulas (e.g. Allen et al. (1998)).
- For winter wheat, the summation in (1) starts on first of January, whereas for spring wheat it starts at sowing, i.e. on the first date of the year when the average temperature over a 10-day period exceeds a threshold  $T_s$ .
- The model sets  $T_b$  equal to 5°C in all cases, but the temperature sum requirements, S, are specified as a function of the long-term mean annual temperature, to reflect the fact that current

- wheat varieties are already chosen to match local climatic settings (Olesen et al. 2012). For similar reasons, the model also expresses the temperature threshold  $T_s$  as function of latitude
- 128 (see Olesen et al. 2012 for details).

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#### 2.5 Probability of heat stress and total precipitation at flowering

- 131 The occurrence of adverse weather conditions in the reproductive period was gauged by
- computing the probability of heat stress  $(p_{HS})$  and total precipitation amounts (tPr) for the 31
- days centred on the date of flowering ( $DOY_{flo}$ ).

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- In a first step, a binary indicator of heat stress occurrence  $(I_{HS})$  was computed for each station
- 136 and day as:

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$$I_{HD} = \begin{cases} 1 & if \ T_d > T_{crit} \\ 0 & otherwise \end{cases}$$
 (3)

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- where  $T_{crit} = 27^{\circ}C$  is the critical temperature (Teixeira et al. 2013), and  $T_d$  is the daytime
- temperature. The latter was estimated from  $T_{min}$  and  $T_{max}$  by integrating in time the equation
- 142 for approximating hourly temperatures derived by Felber et al. (2018), viz.:

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$$T_d = \frac{1}{h_S - h_R - c} \int_{h_R - c}^{h_S} T_h \, dh =$$

$$= T_{min} + (T_{max} - T_{min}) \frac{h_S - h_R + 2a - 2c}{\pi(h_S - h_R - c)} \left\{ 1 - \cos \left[ \frac{\pi(h_S - h_R - c)}{h_S - h_R + 2a - 2c} \right] \right\}$$
(4)

- where  $T_h$  is the temperature at hour h,  $h_S$  and  $h_R$  are sunset and sunrise hours calculated for
- each station based on the station's latitude and the day of the year, a = 2.5 and c = 0.5 (Parton
- 149 & Logan 1981; Felber et al. 2018)

151 With (1) the probability of heat stress was reckoned as:

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$$p_{HS} = \frac{1}{31} \sum_{DOY_{flo}-15}^{DOY_{flo}+15} I_{HD}$$
 (5)

- where  $DOY_{flo}$  is the estimated date of flowering.
- 156 In similar fashion, total precipitation at flowering was evaluated as:

$$tPr = \sum_{DOY_{flo}-15}^{DOY_{flo}+15} Pr$$
 (6)

#### 2.6 Spatial aggregation of the weather and yield data

For the statistical analysis of a possible association between heat stress and drought conditions and wheat yield variability, we spatially aggregated weather and yield data. We identified groups of stations with similar characteristics as of the occurrence of high temperatures by means of cluster analysis. Clustering was implemented following Bador et al. (2015) based on a Partitioning Around Medoids (PAM) approach (Gentle et al. 1991; Bernard et al. 2013) and using the 90th percentile of  $T_{max}$  as a similarity target.

Compared to the more common *K-means* clustering algorithm (Hartigan & Wong 1979;
Hartigan et al. 1981), *PAM* has the advantage of preserving the maxima in each cluster, being
therefore consistent with extreme value theory (Coles 2001). As a measure of distance between
clusters, the F-Madogram was used (Cooley et al. 2006). The analysis was implemented in R
by means of the package SpatialExtremes (Ribatet et al. 2011), with optimal number of clusters
being determined according to an average silhouette coefficient that contrasts cluster tightness
with cluster dissociation.

Results of the analysis indicated two clusters as optimum solution, a Northern Cluster (NC) containing 77 stations essentially located north of 50° N, and a Southern Cluster (SC) with 41 stations located in Southern European Russia and Southern Siberia. By examining the probability of heat stress occurrence in relation to atmospheric blockings, we found it reasonable to further split the NC along the Ural Mountains (approximately  $60^{\circ}E$ ) into a Western (NCW) and Eastern Cluster (NCE, Figure 1).

Although the clustering process primarily targets extreme temperatures, the results are consistent with the Geiger-Köppen climate classification, which displays *Dfa* climates (humid snow climates with hot summer) in the south and *Dfb* climates (humid snow climates with warm summers) in the north (Kottek et al. 2006; Peel et al. 2007). They are also consistent with the the agroclimatic zonation of Bulgakov et al. (2016), which shows dominance of winter wheat cultivation in the *SC*, but of spring wheat cultivation in the *NC*.

Finally, results of the cluster analysis also more or less follow the political division of the Russian territory, with the *NCW* corresponding to large extend to the Central and Volga Federal Districts, the *NCE* to the Ural and the southern part of the Siberian Federal Districts, and the *SC* to the Southern Federal District. Because of this, we used the political borders as limits for the spatial aggregation of the yield data, employing the sown area as a weighting factor for the averaging process.

#### 2.7 Temporal and spatial aggregation of the Blocking Index

To obtain a suitable indicator of blocking activity, we aggregated the 2-dimensional fields of Rohrer *et al.* (2018) over two spatial domains,: the first one spanning the area delimited by  $30 - 60^{\circ}E$  and  $45 - 60^{\circ}N$ , the second the area delimited by  $60 - 90^{\circ}E$  and  $45 - 60^{\circ}N$ .

The size and location of the areas are consistent with those assumed by Schaller et al. (2018) 201 for examining the influence of blocking on Western Russian heatwaves. Here, we considered 202 the first domain as being representative of blocking anticyclones affecting the NCW and SC, 203 the second one as connotative of blocking systems affecting the NCE. 204 205 Analogously to Schaller et al. (2018), we calculated for both domains the area-weighted average 206 of blocked days (ABD), which was taken as blocking indicator for correlation analysis. ABD is 207 the result of spatially averaging the number of positive counts of the BI during the 31 days 208 around flowering  $(DOY_{flo} - 15 \text{ to } DOY_{flo} + 15)$ . 209 210 211 3 Results 212 3.1 Weather conditions at flowering and wheat yields 213 We calculated flowering to occur from end of May to mid-June in winter wheat and from mid-214 June to mid-July in spring wheat (Figure 2). This is in line with the crop calendars published 215 by the US Department of agriculture (United States Department of Agriculture 2006; 2017) and 216 the data extracted from the report prepared by Savin et al. (2007). In the SC flowering occurred 217 earlier than in the NC, especially in the case of spring wheat. Within the NC spring wheat 218 219 flowering in the western part happened a few days earlier than in the eastern part. For winter wheat, the differences in  $DOY_{flo}$  between the NCW and the NCE were minor. 220 221

In general, the SC experienced higher daytime temperatures than either the NCW or the NCE

(supplementary Figure S1), the excess being of about 2°C for both wheat types.

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Total precipitation amounts during the flowering period of winter wheat were of about 56 mm ( $sd = 20.8 \, mm$ ) in the NCW and of 50 mm ( $sd = 18 \, mm$ ) in the NCE and the SC. Corresponding area averaged amounts for the flowering period of spring wheat were higher: namely 71 mm ( $sd = 26.6 \, mm$ ), 62 mm ( $sd = 15.9 \, mm$ ) and 56 mm ( $sd = 18.9 \, mm$ ) in the NCW, NCE and SC, respectively.

Concerning yield, spatial aggregation of the provincial statistics indicated area-mean productivity levels of 22.0, 18.8 and 26.1  $dt ha^{-1}$  for winter wheat, and of 17.8, 13.8 and 9.1  $dt ha^{-1}$  for spring wheat, in the *NCW*, *NCE* and *SC*, respectively (Table 1). For both wheat types and all clusters, the within-cluster coefficient of variation was approximately 25%, in line with the findings of Zampieri et al. (2017).

At the annual level, largest yield deficits occurred in 1998, varying, depending on region, between -30 and -60% in the case of winter wheat, and between -40 and -80% in the case of spring wheat (Table 1). Yields were again low in 2010, falling short of the average by more than 30% in the case of winter wheat, and by more than 40% in the case of spring wheat in the west of the study area (*NCW* and *SC*). The results suggest that the 2010 heat wave had only minor effects on wheat production to the east of the Ural Mountains.

Table 1: Yield statistics valid for 1995-2014, and yield anomalies in 1998 and 2010.

| yield 1995-2014                     | 7     | winter wheat |       | spring wheat |       |       |  |
|-------------------------------------|-------|--------------|-------|--------------|-------|-------|--|
| 7                                   | NCW   | NCE          | SC    | NCW          | NCE   | SC    |  |
| mean (dt ha <sup>-1</sup> )         | 22.0  | 18.8         | 26.1  | 17.8         | 13.8  | 9.1   |  |
| st. dev. (dt ha <sup>-1</sup> )     | 6.3   | 5.3          | 7.1   | 4.8          | 3.7   | 2.4   |  |
| Coefficient of variation            | 28.7% | 28.1%        | 27.1% | 26.8%        | 26.8% | 26.1% |  |
| anomaly 1998 (dt ha <sup>-1</sup> ) | -7.68 | -10.68       | -7.44 | -10.3        | -5.05 | -7.07 |  |
| anomaly 2010 (dt ha <sup>-1</sup> ) |       |              |       |              |       |       |  |

|     | -7.37 -5.06 -5.56 -7.92 -0.95 -4.15  |
|-----|--|
| 245 | -7.57 -5.00 -5.50 -7.52 -0.55 -4.15  |
| 246 | 3.2 Probability of heat stress, precipitation and their association to yield variability           |
| 247 | Our analysis indicates that sites frequently exposed to high temperatures were mostly located      |
| 248 | in the SC (supplementary Figure S2). The probability of heat stress was higher in spring than      |
| 249 | winter wheat owing to differences in the time of development (Figure 2).                           |
| 250 |  |
| 251 | Regional differences also appear in relation to total precipitation amounts for the 31-day         |
| 252 | window around the flowering date (supplementary Figure S3). According to our evaluation,           |
| 253 | stations with the highest probability of heat stress also displayed the lowest total precipitation |
| 254 | amounts. $NC$ stations received more precipitation than $SC$ stations and precipitation amounts    |
| 255 | were slightly higher for spring than for winter wheat.   |
| 256 |  |
| 257 | Results of the correlation analysis of yields against probability of heat stress and total         |
| 258 | precipitation amounts are presented in Table 2 and in Figures 3 and 4. In view of the relatively   |
| 259 | small sample size (time series of yield span only 20 years) and because data cannot be assumed     |
| 260 | to stem from a bivariate normal distribution, we quantified the level of association by means of   |
| 261 | Spearman rank correlation, a more robust metrics that the ordinary Pearson correlation             |
| 262 | coefficient (Press et al. 2007).   |
| 263 |  |
| 264 | Heat stress probability in the reproductive period did not explain winter wheat yield variations   |
| 265 | in the NCW and the SC. We found however, a significant negative correlation between yield          |
| 266 | and probability of heat stress for winter wheat in the NCE and for spring wheat in all clusters.   |
| 267 | Regarding the association of wheat yield variations with total precipitation, we obtained          |
| 268 | significant positive correlations for both winter and spring wheat in the NCE and for winter       |

wheat in the SC. Results were not significant for the NCW.

For winter wheat, there was a significant correlation between probability of heat stress occurrence and total precipitation amounts in the *NCW* and *NCE* and for spring wheat, in the *NCW* and *SC* (supplementary Table S1).

Table 2: Spearman rank correlation between probability of heat stress, respectively total precipitation and wheat yields. Bold face: statistically significant results. Significance levels: (\*) 10%, (\*\*) 5%, (\*\*\*) 1%. Results for the one-side test of a negative ( $p_{HS}$ ), respectively positive (tPr) association against the null hypothesis of no association.

|          | Winter Wheat |             |           |   | Spring Wheat |            |            |  |
|----------|--------------|-------------|-----------|---|--------------|------------|------------|--|
|          | NCW          | NCE         | SC        |   | NCW          | NCE        | SC         |  |
| $p_{HS}$ | -0.02        | -0.72 (***) | -0.28     | 1 | -0.47 (**)   | -0.56 (**) | -0.47 (**) |  |
| tPr      | 0.31         | 0.68 (***)  | 0.53 (**) |   | 0.24         | 0.46 (**)  | 0.37       |  |

The lack of correlation between probability of heat stress at flowering and winter wheat yield variations in the *NCW* can be understood bearing in mind the early date of flowering and the relatively low temperatures occurring in this cluster at this time of the year. The lack of correlation between winter wheat yield variations and occurrence of heat stress in the *SC* (Figure 3) is less obvious and further discussed in Section 4. A priori, however, numerical problems arising from uncertainties in the specification of the critical temperature thresholds (which depends on wheat variety) or in the estimation of the date of flowering cannot be excluded.

To ascertain a possible impact of uncertainties in the date of flowering, we carried out a simple sensitivity analysis, that consisted in subtracting 5 days and adding 5 and 10 days to the estimate obtained with the Olesen et al. (2012) model and reassessing the statistics. We found that shifting the date of flowering had minor effects on the correlation coefficients calculated for

spring wheat (all cluster) as well as for winter wheat in the *NCW* and *NCE* (supplementary Table S2). However, it affected the correlation coefficient and its statistical significance in the case of winter wheat in the *SC*. In this case, the assumption of later dates of flowering resulted in a higher correlation coefficient.

## 3.3 Links to blocking

The time series plotted in Figure 5 and supplementary Figures S4 and S5 demonstrate the distinct exposure to blocking conditions in the different clusters. Long-term mean *ABD* values ranged from about 1 day in the *NCE* to about 3 days in the other two clusters. In the *NCW* and *SC*, the average number of blocked days was slightly higher during the flowering season of *WW* than during that of *SW*. The inter-annual variability of the average number of blocked days was large in all cases. In the *NCW* and the *SC*, the standard deviation of *ABD* was around 2 days for winter wheat, and around 1 day for spring wheat. For the *NCE*, the standard deviation was around 1 day for both wheat types. Overall, these results reflect the different dynamical and climatological characteristics of blocking situations occurring to the east and to the west of the Ural mountains (Matsueda 2011; Cheung et al. 2013; Dunn-Sigouin & Son 2013).

With respect to winter wheat in the SC, and to spring wheat in the NCW and the SC, the variability in blocking activity explained about 40% of the inter-annual variability in probability of heat stress at flowering (supplementary Figure S4 and Table 3). Blocking was also largely responsible for precipitation variations in the NCW (supplementary Figure S5 and Table 3). In this case, the share of explained variance was less important in the NCE (around 20% for both wheat types) and the SC (14% in WW and 24% in SW).

Table 3: Spearman rank correlation between average number of blocked days and (top to bottom)
heat stress, total precipitation and weighted yield. Bold face: statistically significant results.

Significance levels: (\*) 10%, (\*\*) 5%, (\*\*\*) 1%.

|                | Winter wheat |             |            |  | Spring wheat |             |             |  |
|----------------|--------------|-------------|------------|--|--------------|-------------|-------------|--|
|                | NCW          | NCE         | SC         |  | NCW          | NCE         | SC          |  |
| $p_{HS}$       | 0.53 (***)   | 0.44 (***)  | 0.62 (***) |  | 0.74 (***)   | 0.43 (***)  | 0.76 (***)  |  |
| tPr            | -0.73 (***)  | -0.48 (***) | -0.38 (**) |  | -0.79 (***)  | -0.42 (***) | -0.49 (***) |  |
| Weighted yield | -0.33        | -0.36       | -0.37      |  | -0.54 (**)   | -0.36       | -0.57(***)  |  |

For spring wheat, year-to-year variations in the average number of blocked days explained around 30% of the observed yield variations in the *NCW* and *SC* and 13% in the *NCE* (Figure 5 & Table 3). For winter wheat, the correlation coefficients were lower and overall statistically not significant.

An interesting feature emerging from Figure 5 is the fact that, except for spring wheat in both the *NCW* and the *SC*, the average number of blocked days estimated for 2010 was close or even below the long-term average. This contrasts with the situation in 1998, a year for which the average number of blocked days fell in the upper tail of the distribution for both wheat types and all clusters. One possible explanation for the unexpected result concerning 2010 was the geographic position of the atmospheric blocking systems during spring season (Dole et al. 2011). In fact, it was only during June, i.e. the flowering time of *SW* that persistent high-pressure centres settled close to the western border of the study area.

#### 4 Discussion and Conclusion

The sensitivity of wheat to extreme high temperature and water shortage around flowering makes wheat cultivation particularly vulnerable to heat waves that occur at this stage of development (Wheeler et al. 2000; Gourdji et al. 2013). A first goal of our study was to quantify the correlation between the probability of heat stress occurrence and total precipitation in the

reproductive period with yield, based on regional yield statistics. A second goal was to investigate the role of atmospheric blocking as a determinant of extreme weather conditions during flowering and eventually also with yield.

Because we were not able to obtain actual phenological data, we employed the phenology model of Olesen et al. (2012) to estimate the dates of flowering. We showed that the model output is in line with agronomic information available from independent sources, for example the crop calendar published for the Russian Federation by the International Production Assessment Division of the Foreign Agricultural Service of the USDA (United States Department of Agriculture 2017). Lack of phenological data from Russia precluded a direct model validation, implying that, our estimates are uncertain. In view of this, we recommend the collection of historical and current phenological data from all over Russian arable areas as a primary task for the future.

Our work confirms the crucial role of heat stress and water shortage for Russian wheat production, supporting the conclusion of Zampieri et al. (2017) that heat stress can explain a large fraction of the variability in winter wheat yields observed for Russia at the national scale. However, the correlation was significant only for the *NCE*, which suggests that other determinants are need to be taken into account in the other regions.

For spring wheat, on the other hand, we found significant correlations between heat stress and yield for all regions. In spring wheat, flowering occurs later in the year, and temperatures are generally higher than during the flowering period of winter wheat. We therefore expect higher levels of heat stress and a higher incidence on yield in spring wheat than winter wheat.

Concerning the link between precipitation amounts at flowering and yield, significant high correlations were established in the *SC* and the *NCE* for both wheat types, reflecting the fact that aridity patterns are more persistent in time (and space) and therefore more coherent between winter and spring wheat (Fig. S3).

We ascertained that the probability of heat stress is particularly important in the Southern and Central Federal Districts. The former region is the most important cultivation area with respect to winter wheat; the second contributes significantly to spring wheat production. It is also important to bear in mind that heat stress and water shortage usually occur in sympathy, which magnifies the negative impacts of these two types of stress taken individually (Nicolas et al. 1984). Indeed, correlation coefficients between heat stress and water shortage were in many cases significant (Table S1). Years characterized by a high probability of heat stress and low precipitation in the present analysis (in particular 1981, 1998 and 2010) are also identified in the compilation of historical drought events and heat waves prepared for Eurasia by Schubert et al. (2014).

For Russia, previous studies proposed atmospheric blocking as a precursor of extreme weather (Grumm 2011; Schneidereit et al. 2012; Antokhina et al. 2016). The role of blocking as a driver of warm spell is in fact well established (Pfahl & Wernli 2012; Brunner et al. 2017; Schaller et al. 2018). Our analysis confirmed that for Russia's wheat production area, blocking activity had a systematic influence on the occurrence of adverse weather conditions at flowering, explaining between 20 to 60% of the inter-annual variability of heat stress and water availability at flowering in the time since 1980.

The results suggest regional differences between the impacts of European blocking systems on wheat production in the north-west and south of Russia, and the repercussions of Ural blocking

systems on wheat production in the northeast. We found in particular a significant negative correlation of around -0.6 between the spring wheat variability and blocking activity for the *NCW*, *SC* and -0.4 for the *NCE*.

The lack of correlation between blocking activity and probability of heat stress with respect to winter wheat production in Southern Russia is in any case surprising. Apart from possible problems in assessing the date of flowering (see above), the absence of a negative association could also reflect the rapid expansion of the cultivated area that has taken place in the last two decades and efforts undertaken to intensify winter wheat production in the areas offering the highest production potentials (Schierhorn et al. 2014). It could also reflect the fact that also frost and heavy rains are often the cause of damages. Frost and heavy rains have been recurrent in the recent past in southern Russia, as documented e.g. in the Commodity Intelligence Report on the Russian Federation by the International Production Assessment Division (IPAD) of the Foreign Agricultural Service of the USDA (United States Department of Agriculture 2018). An interesting future study would therefore be to examine the role of blocking during winter and early spring season, connected with cold spells (Sillmann et al. 2011; Brunner et al. 2017), taking into account key agronomic and production data.

A chief premise of our study was that wheat is particularly sensitive to heat stress during flowering (Brouwer et al. 1989; Wheeler et al. 2000; Wollenweber et al. 2003). Implicitly we assumed that limiting the attention to a short period (31 days around flowering) was adequate for the purpose of uncovering statistically significant relations. It is possible however, that it would have been more suitable to consider the entire growing season. To address this, we repeated the analysis aggregating all relevant data over days of the year 121-212 for *WW* (May-Jul); and days 152-243 for *SW* (Jun-Aug, http://www.amis-outlook.org/amis-about/calendars/en/). Correlations were neither higher nor statistically more significant than we

found with periods used in the main analysis. However, we used the 31-days window around 417 blooming to be consistent with previous assessments (Teixeira et al. 2013; Gourdji et al. 2013). 418 419 Russian wheat production could benefit from early warning systems that help limiting the 420 negative impacts of extreme weather on crop development. A better physical understanding of 421 the link between atmospheric blocking and extreme heat would certainly support such a system. 422 This requires addressing teleconnection patterns as precursors of blocking situations, on the one 423 hand, and heat-waves and droughts on the other hand (Renwick & Wallace 1996; Shabbar et 424 al. 2001; Barriopedro et al. 2006; Rust et al. 2014; Folland et al. 2009; Casanueva et al. 2014; 425 426 Scherrer et al. 2006). 427 During the summer season teleconnection patterns are weaker but they have been proven to 428 have strong links with temperature and precipitation conditions in Eurasia (Rust et al. 2014; 429 Folland et al. 2009; Irannezhad & Kløve 2015; Irannezhad et al. 2016). Other studies 430 demonstrated that different phases of El Niño-Southern Oscillation (which is usually perceived 431 more clearly during boreal winter) are statistically related to spring and summer temperature 432 and precipitation anomalies in Russia (Mokhov & Timazhev 2017; Mokhov & Semenov 2016). 433 434 Further work is hence required to clarify the interdependence of blocking activity, with largescale circulation anomalies and with El Niño phases in the areas that support Russian wheat 435 production. 436 For the present investigation, we opted for a binary blocking index that quantifies blocking 437 occurrence without addressing the multifaceted nature of blocking dynamics including 438 intensity, location, and life-cycle of individual blocking events. Additional work addressing 439 these aspects could help explaining the causal relations leading to adverse conditions for wheat 440 cultivation, eventually reducing the vulnerability of Russian agriculture to extreme weather 441 442 events.

In principle, expanding production areas to the North (Fischer et al. 2002; Schierhorn, 443 Faramarzi, et al. 2014; Liefert & Liefert 2015; Di Paola et al. 2018; Belyaeva & Bokusheva 444 2018) or changing wheat types and/or varieties could be a measure to avoid production 445 shortfalls caused by critically high temperatures. Yet, in practice, the risk of incurring in 446 production shortfalls caused by heat stress would remain substantial (Dronin & Kirilenko 447 2008). 448 449 Looking forward, an important question for Russian agriculture is the one of a possible future 450 increase in blocking activity. Mokhov & Timazhev (2015) studied the likelihood of future 451 occurrences of northern hemisphere blocking episodes from the perspective of different 452 emission and climate scenarios and concluded that a tendency for an increase in the duration of 453 blocking events is very likely. 454 455 Conclusions of retrospective studies are, however, less categorical in this respect. Barnes et al. 456 (2014) found for example that no clear increase in blocking could be detected over 1980-2012 457 irrespective of the data and indices used. Rather they noticed that blocking exhibits large 458 variations and that it is difficult to discriminate between the contribution from internal 459 variability and external forcing. Candidates for the latter are e.g. a weakening summer 460 circulation in the Northern Hemisphere mid-latitudes (Coumou et al. 2015) or changes in the 461 Atlantic Multidecadal Ocean Variability (Häkkinen et al. 2011). 462 463 Consistent with this, Dunn-Sigouin & Son (2013) concluded that even under the constraints of 464 the RCP 8.5 emission scenarios, expected changes in the number and duration of blocking 465 events under future climatic conditions will likely be insignificant. Clearly, there is still 466 considerable uncertainty in projections of blocking activity under future climate scenarios 467 (Woollings et al. 2014), which stresses the need for additional investigations.

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#### 720 {Bibliography

#### 721 Figures

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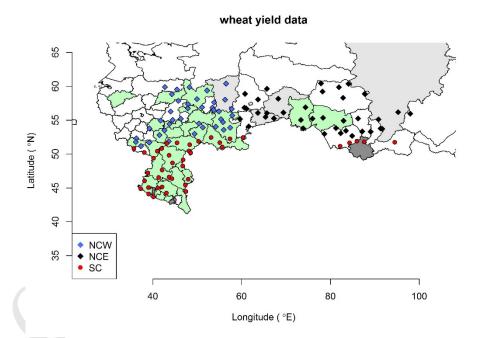
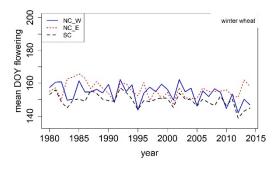


Figure 1: Map of the study area showing the political borders (oblasts, i.e. subjects of the Russian Federation), the location of the weather stations, grouped according to the results of the cluster analysis (blue diamonds: *NCW*; black diamonds: *NCE*; red dots: *SC*). In addition, the maps indicate whether both types of wheat are cultivated (light green), or whether there is a predominance of either winter (dark grey) or spring wheat (light grey).



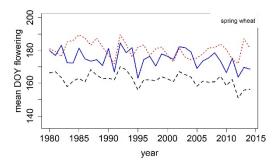
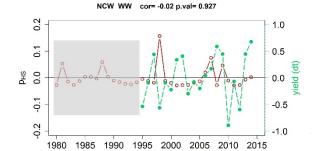
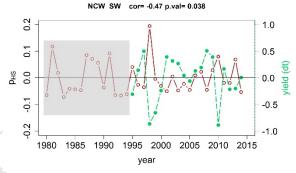
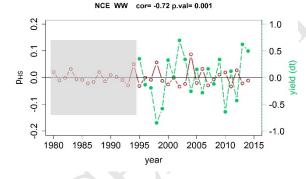
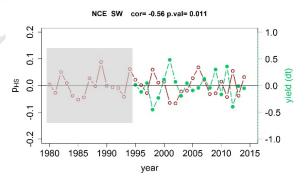


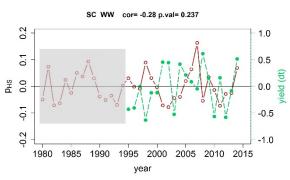
Figure 2: Time series (1980-2014) of the mean date of flowering (day of the year,  $DOY_{flo}$ ) for WW and SW, spatially averaged for each cluster separately.











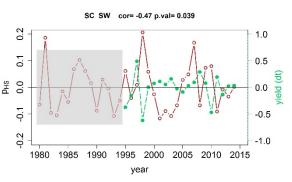


Figure 3: Time series of de-trended mean probability of heat stress (red continuous lines and open dots) and de-trended mean wheat yield (green dashed lines and solid dots) in the reproductive period of *WW*, respectively *SW*. Shown from top to bottom are the series for the *NCW*, *NCE* and *SC*. Correlation coefficient and statistical significance are indicated above each panel. Grey boxes mask the years for which wheat yield data was not available.

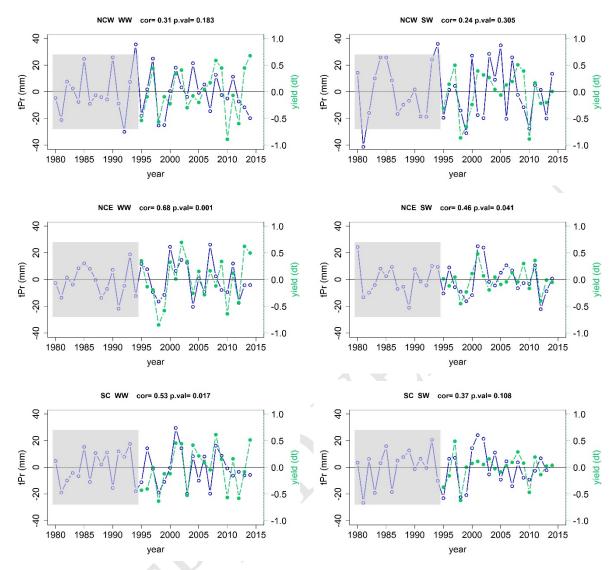
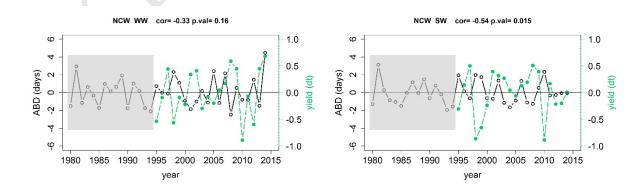


Figure 4: Same as Figure 3 but for de-trended mean precipitation amount (blue lines and open dots) instead of probability of heat stress occurrence.



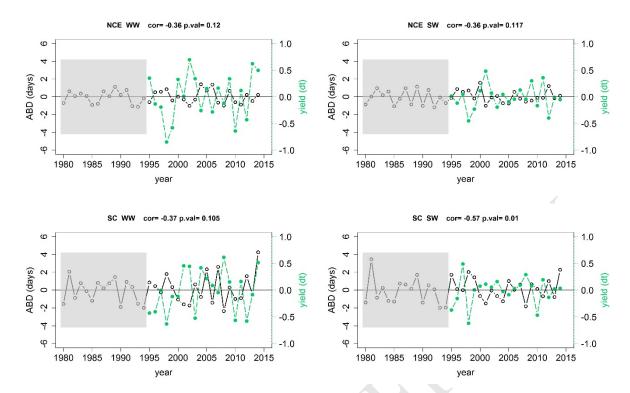
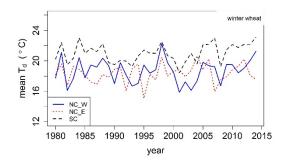


Figure 5: Same as Figure 3 but contrasting the time series of yield (green dashed lines and solid dots) with the time series of average number of blocked days (black continuous lines and open dots).

## **Appendix**



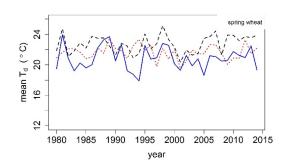


Figure S1: Time series (1980-2014) of the mean daytime temperature ( $T_{day}$ ) for WW and SW. The data represent mean values over the 31 days around flowering ( $DOY_{flo} - 15$  to  $DOY_{flo} + 15$ ), spatially averaged for each cluster separately.

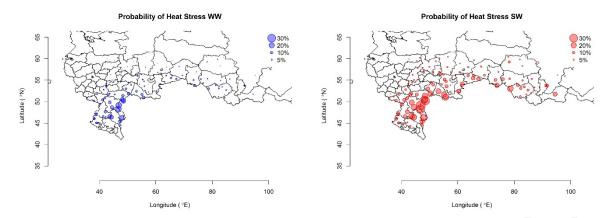


Figure S2: Mean (10980-2014) probability of heat stress occurrence in the 31 days around flowering for *WW* and *SW*. The size of the dots is is proportional to the probability.

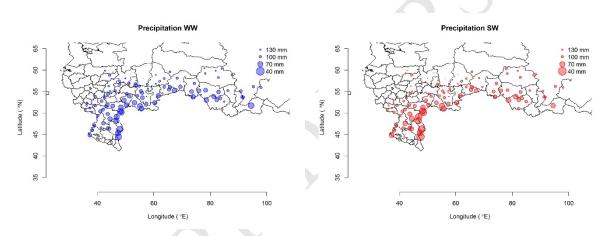
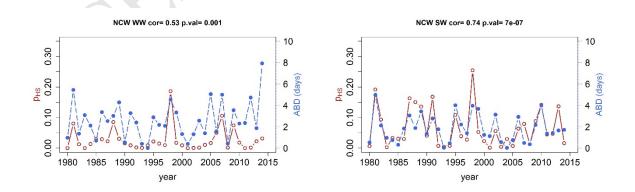


Figure S3: Same as Figure S2 but for the mean total precipitation amount. Note that in this case the size of the dots is inversely proportional to precipitation volume.



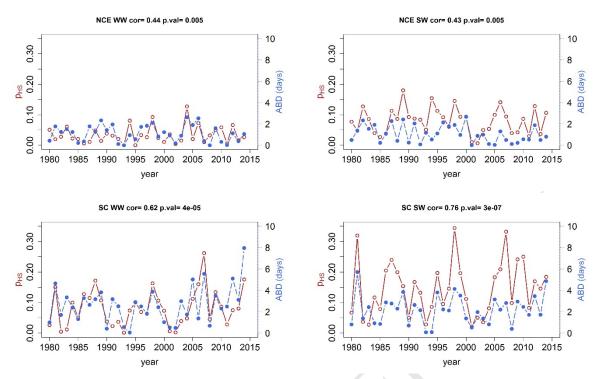
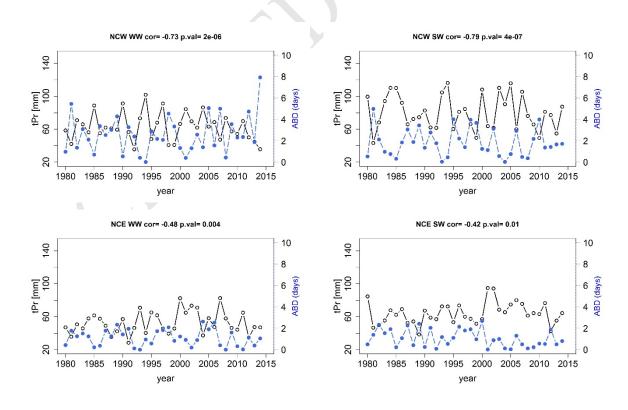


Figure S4: Time series (1980-2014) of the mean *ABD* (blue dashed line and solid dots) and the mean probability of heat stress (red continuous line and open dots) 31 days around *WW* and *SW* flowering for *NCW*, *NCE* and *SC*. Correlation coefficient and p-value are indicated above each panel.



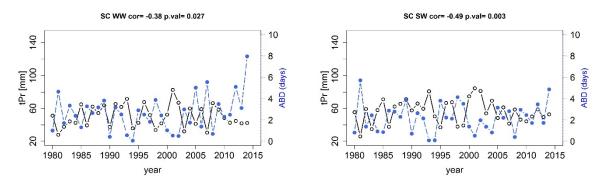


Figure S5: same as Figure S4 but with time series of total precipitation amounts (black continuous lines and open dots) instead of the probability of heat stress occurrence.

Table S1: Spearman rank correlation between probability of heat stress and mean total precipitation amounts in the 31 days around WW and SW flowering. Bold face: statistically significant results. Significance levels: (\*) 10%, (\*\*) 5%, (\*\*\*) 1%.

| Winter wheat |             |          |             | Spring wheat |            |
|--------------|-------------|----------|-------------|--------------|------------|
| NCW          | NCE         | SC       | NCW         | NCE          | SC         |
| -0.44 (**)   | -0.62 (***) | -0.4 (*) | -0.63 (***) | -0.35 (*)    | -0.47 (**) |

Table S2: Sensitivity analysis of the correlation between heat stress and yield variations with respect to shifts in the date of flowering. Bold face: statistically significant results. Significance levels: (\*) 10%, (\*\*) 5%, (\*\*\*) 1%.

| NCE 2 -0.65 (**) | SC -0.04  | NCW<br>-0.48 (*) | NCE -0.48 (*)                         | SC -0.49 (*) |
|------------------|-----------|------------------|---------------------------------------|--------------|
| · · ·            | -0.04     | -0.48 (*)        | -0.48 (*)                             | -0.49 (*)    |
| 2 2 72 (11)      |           |                  |                                       |              |
| 2 -0.72 (**)     | -0.28     | -0.47 (*)        | -0.56 (*)                             | -0.47 (*)    |
| 3 -0.66 (**)     | -0.31 (*) | -0.46 (*)        | -0.69 (**)                            | -0.53 (**)   |
| 6 -0.59 (**)     | -0.37 (*) | -0.51 (*)        | -0.63 (**)                            | -0.55 (**)   |
|                  | ` ′       |                  | · · · · · · · · · · · · · · · · · · · |              |

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We assessed the plausibility of the predicted sowing dates for spring wheat by comparing them to data extracted from Savin et al. (2007) and the range of dates given in Agricultural Market

Information System (AMIS, 2012), United States Department of Agriculture 2006, 2017, Figure S6).

As for the dates of flowering, we compared the estimated dates with the potential heading periods given in United States Department of Agriculture (2006, 2017, Figures S7, S8). According to Acevedo et al. (2002), heading precedes flowering by approximately 10 days. Given that the ranges of possible heading dates indicated in the reports by the United States Department of Agriculture (2006, 2017) extend over more than one month, we simply took these as proxies for the ranges of possible flowering dates.

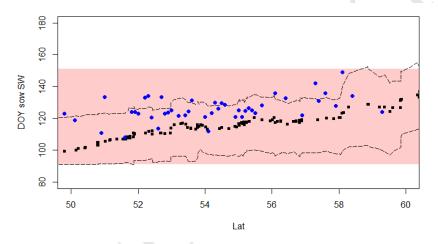


Figure S6: Predicted mean sowing dates for spring wheat (black dots) and range (dashed lines) against actual dates extracted from Savin et al. (2007) (blue dots). The figure also shows in the background (red stripe) the range of heading dates given in United States Department of Agriculture (2006, 2017).

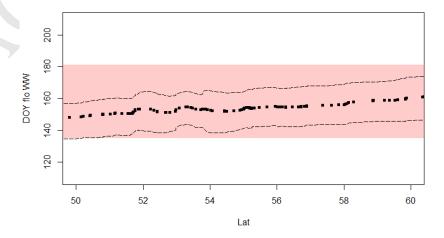


Figure S7: Computed range (dashed lines) and mean dates of flowering (dots) for WW as a function of latitude. The figure shows in the background (red strip) the range of heading dates given in United States Department of Agriculture (2006, 2017).

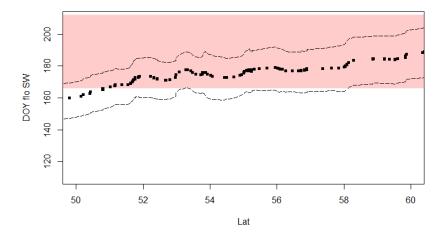


Figure S8: Same as Figure S7, but for SW.

 As seen in the figures, predictions by the Olesen et al. (2012) model are overall realistic and in line with independent information. The only exception are the flowering dates of *SW* in the very south of the study area. In Southern Russia, winter wheat is the predominant type. It is likely that the overall range of dates available from United States Department of Agriculture (2006, 2017, red stripe Figure S8) is biased toward information available from more northern latitudes.